

The effect of globalization, income and tourism on environment: An empirical analysis

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Abstract

This study investigates the impact of globalization, real income, and tourism on the environment in top 10 destinations that attract the most tourists, using panel unit root, panel cointegration, and panel cointegration estimators for the period between 1995 and 2014. Panel cointegration test results show that the series moves together in the long run. According to the long-run panel estimator results, an increase in real income negatively affects the environment in China, France, Spain, Thailand, and the UK. Tourism decreases environmental degradation in Germany, Italy, the UK, and the United States. Globalization reduces environmental degradation in France and the UK. The short-run panel estimator shows that an increase in real income in all countries, except China, increases environmental degradation. Tourism decreases environmental degradation in Mexico and increases it in Spain. Finally, globalization contributes to the reduction of environmental degradation in Italy.

Keywords: globalization, tourism, real income, environmental degradation, panel data analysis

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1. Introduction

The environment plays a central role in maintaining uninterrupted and healthy functioning of daily life activities (Hosseini and Soltani, 2011; Sun et al., 2018). However, environmental quality degradation, due to social and economic activities, has started to seriously threaten all areas of life (Sani et al., 2020; Shengjuan and Jingping, 2011; Warford and Partow, 1989). The United Nations defines environmental degradation as 'deterioration in environmental quality from ambient concentrations of pollutants and other activities and processes, such as improper land use and natural disasters' (United Nations, 1997). Environmental degradation stems from industrial production and consumption activities (Nasir et al., 2021; Ozcan et al., 2020; Sharif et al., 2020a). In this context, the industrial revolution was a turning point in accelerating this degradation (Burnete and Choomta, 2015; Maneejuk et al., 2020; Munir and Riaz, 2020). Carbon dioxide (CO₂) emissions are one of the main indicators of environmental degradation (Alvarado et al., 2018; Asongu and Odhiambo, 2020; Ozcan et al., 2020; Saud et al., 2020; Wang and Dong, 2019). Before the industrialization process, CO₂ emissions in the world were much lower than what they are today (Kim and Park, 2020; Moerschbaecher and Day, 2010). Increasing fossil fuel consumption along with capital-intensive production techniques has increased the relative share of toxic gases in the air, such as CO₂ (Anwar et al., 2020; Saud et al., 2020; Tsou and Edelman, 2013; Waqih et al., 2019).

The industrial revolution produced significant quantitative and qualitative changes in production and accelerated the globalization process as well (Ateş, 2008; Sachs, 2017). Gurgul and Lach (2014) define globalization as the process of goods and capital market unification around the world, in which barriers to international trade and foreign investment are reduced. The removal of barriers between countries connects economies and results in various consequences in all areas of life (Ayomitunde et al., 2020; Radulović and Kostić, 2020; Shahbaz et al., 2015; Wang et al., 2019). One of the variables in globalization is environmental conditions (Baek et al., 2009; Shahbaz et al., 2015; Yang et al., 2019).

There are viewpoints that explain the impact of globalization on the environment in different ways. The first of these viewpoints is the pollution haven hypothesis (PHH). According to this hypothesis, producers who want to minimize their costs transfer their economic activities to countries with weak environmental standards, through foreign direct investment (FDI) (Candau and Dienesch, 2017; Dong et al., 2012; Shao, 2018; Zugravu-Soilita, 2017). PHH theory posits that polluting industries move from countries with strict environmental regulations to countries with weak regulations (Ayadi et al., 2019; Cole, 2004; Dinda, 2004; Guzel and Okumus, 2020; Mert and Caglar, 2020). However, the pollution halo hypothesis argues that FDI has a positive effect on the environment (Demena and Afesorgbor, 2020; Jun et al., 2018; Repkine and Min, 2020). The halo effect hypothesis states that foreign-funded companies spread knowledge and that domestic businesses implement environmentally friendly practices, while improving their environmental performance (Doytch and Uctum, 2016; Mabey and McNally, 1999; Mert and Caglar, 2020; Repkine and Min, 2020; Singhania and Saini, 2021).

Simultaneously, globalization alters the environmental balance through three channels, namely the scale effect, the composition effect, and the technique effect (Grossman and Krueger, 1991; Rafindadi and Usman, 2019; Rahman, 2020; Sethi et al., 2020; Shahbaz et al., 2015). According to the scale effect, economic activities and total production levels expand because of trade liberalization. The expansion of economic activities significantly increases the use of natural resources and causes environmental degradation (Aktaş, 2015; Etokakpan et al., 2020; McAusland, 2008; Saidi and Hammami, 2015; Shahbaz et al., 2016). The globalization process also creates changes in the mix of production and consumption activities (Copeland, 2009). The transition from the agricultural sector to the industrial sector may lead to higher energy consumption and CO₂ emissions, but the transition from the industrial sector to the

service sector is expected to reduce emissions (Aller et al., 2015; Frankel, 2003; Grossman and Krueger, 1995; Zheng et al., 2015). The technical effect implies a reduction in pollution emissions due to increased demand for stricter environmental regulations because of economic growth (Honma, 2015; Saud et al., 2020; Shahbaz et al., 2016). Generally, it can be said that the scale effect has negative consequences on environmental conditions, but the technical effect can reduce this negative effect. The environmental consequences of the composition effect may vary (Tamiotti et al, 2009).

Integration into the global economic system has also significantly changed countries' income levels. The relationship between income and environmental degradation is explained within the framework of the Environmental Kuznets Curve (EKC) hypothesis (Ahmed and Long, 2012; Amar, 2021; Ozcan et al., 2020). This hypothesis states that environmental degradation increases in the early stages of economic growth but economic growth leads to environmental improvement, after reaching a certain per capita income level (Özmen et al., 2019; Rehman and Rashid, 2017). After reaching a certain income level, education levels rise and environmental awareness becomes stronger as the economy transitions from industrial based to service sector based (Dinda, 2004; Isik et al., 2019; Mazur et al., 2015; Panayotou, 1993).

The tourism sector is one of the indicators whose effects on the environment are disputed. In this study, the relationship between tourism and the environment is discussed in detail, both theoretically and empirically. This study contributes to other studies in the literature that examine the effects of globalization, income, and tourism on the environment. National and international tourism's contribution to air pollution was first discussed theoretically by Bach and Gößling (1996). They stated that the aviation industry increases greenhouse gas emissions, such as CO₂ (Folke et al., 2006; Koçak et al., 2020; Scott et al., 2012; Stefan, 2010). Following this research, interest in the relationship between tourism and CO₂ emissions increased. With the effects of globalization and economic growth, the number of tourists in the world is increasing day by day, so the environmental dimensions of tourism have become critical.

According to the (UNWTO and ITF, 2019), international tourist arrivals, which were 770 million in 2005, increased to 1.2 billion in 2016 and are estimated to be 1.8 billion in 2030. Domestic tourist arrivals are expected to reach 15.6 billion in 2030 from 4 billion in 2005. Since the transportation sector is one of the main elements in the tourism sector, this increase in tourism-related activities increases energy consumption and CO₂ emissions. Therefore, tourist arrivals and departures increase energy-related CO₂ emissions (Amin et al., 2020; Dogan and Aslan, 2017; Liu et al., 2019). International tourism revenues positively affect economic growth by increasing net exports (Aratuo et al., 2019; Brida et al., 2020; Eyuboglu and Uzar, 2020; Leitão and Lorente, 2020; Selvanathan et al., 2020). As economic growth stimulates consumption and investment expenditures, energy demand expands. All tourist activities are based on energy from fossil fuels that contribute to CO₂ emissions (Balli et al., 2019; Jebli et al., 2019; Mikayilov et al., 2018; Nepal et al., 2019; Zhang et al., 2019). Although the negative effects on the environment are emphasized here, tourism may also affect the environment positively, because economic growth caused by increasing tourism revenues improves living standards (Dayananda, 2016). Within the framework of these explanations, we will investigate whether globalization, income, and international tourism activities impact CO₂ emissions by analyzing data from the period between 1995 and 2014 in the 10 countries that attracted most tourists, through panel long-run FMOLS and DOLS and panel short-run PMG and MG coefficient estimation methods. Our study was guided by Akadiri et al. (2019c), Akadiri et al. (2019b) and Akadiri et al. (2020b) whose studies examined the causal link between CO₂ emissions, globalization, economic growth, and tourism. However, these studies are specific to the 10 top tourist destinations. The essence of this study, considering the top 10 countries, stems not only

from the countries' strong tourism activities, but also from the countries' structural economic characteristics. Therefore, this study examines the short- and long-run relationship between variables (CO₂ emissions as a measure of pollutant emissions, real income, globalization, and tourism) observed in the top 10 tourist-attracting countries. This study adds novelty to the literature, potentially presenting a unique difference from previous studies. Given the uniqueness of the sample economies, this is the first study to support the 'myth' of short- and long-term. Because it had previously been unclear if there was a significant difference between the two periods in the economies of the top 10 countries. Also, to the knowledge of the authors, this is the first study to examine the environmental impact of globalization on the top 10 countries.

The rest of this paper is structured as follows. Section 2 focuses on the literature review. Section 3 describes the model, methodology, and data. Section 4 reports and discusses the empirical findings and finally, Section 5 concludes with some policy implications.

2. Literature Review

In the literature, there are a number of studies addressing the impact of globalization, income, and tourism on the environment. However, there are only a limited number of studies that examine the relationship between these four variables together. To begin with, we examine studies on the relationship between globalization and the environment, income and the environment, and tourism and the environment separately. Then, we briefly present the information about the studies we could access that cover these four variables.

Though there are many studies test the relationship between income and environment, most recent studies emphasize causality. While the samples and analysis periods in these studies were different, results indicate that the direction of causation between the two variables was also different. Magazzino (2016) examined the period between 1960 and 2013 in Gulf Cooperation Countries (GCC). According to the causality test findings, income causes environmental pollution in Qatar, Oman, and Kuwait. Ahmad et al. (2019)'s results from a sample of Chinese provinces show that increases in income, increase environmental pollution. The results of Khan et al. (2020) show that economic growth increases the CO₂ emissions for the period between 1965 and 2015 in Pakistan. In contrast, Gokmenoglu and Sadeghieh (2019) found that income increases have a positive effect on the environment in the period between 1960 and 2011 in Turkey. Chen et al. (2019) conclude, with the data from 1980 to 2014, according to causality test findings, there is a bidirectional causality between income and environment in China. Likewise, Acheampong (2018) finds bidirectional causality between income and environment in Caribbean-Latin America areas. Banday and Aneja (2019) indicate causality from income to environmental pollution in 4 G7 countries, and from environmental pollution to income in a single country. Gövdeli (2019) identifies causality from income to environmental pollution in 26 OECD countries and Stamatiou and Dritsakis (2019) determine unidirectional causality from income to environmental pollution in Italy. The tests performed by Zaidi and Ferhi (2019) show a bidirectional linear association between income and the environment in sub-Saharan countries. Adedoyin et al. (2020) demonstrate that in the case of BRICS countries, there is a positive interaction between income and the environment. However, there are studies in which causality could not be determined. In this context, Cai et al. (2018), Gorus and Aydin (2019), and Salahuddin and Gow (2014) indicate that there is no causal relationship between income and the environment.

Socio-cultural needs, brought about by increases in income, accelerate tourism activities worldwide. Tourism revenues are an important part of economic growth in many countries. In this context, the direction of the link between tourism and environmental balance has begun to be widely considered in

empirical studies. While Lee and Brahmašreṇe (2013) conclude that tourism in the European Union reduces environmental pollution, Alam and Paramati (2017), Anser et al. (2020) and Sharif et al. (2020b) find that tourism is the cause of environmental pollution in some selected countries. Dogan and Aslan (2017) evidence unidirectional causality from tourism to environmental pollution in the European Union and candidate countries. Whereas, Paramati et al. (2017b) show that in countries located in the western part of the European Union, tourism reduces environmental pollution, and increases it in countries located in the eastern part. However, Paramati et al. (2017a) in another study, find bidirectional causality in developed countries but did not find causal relationships between tourism and the environment in developing countries. Likewise, Satrović and Muslija (2019) also did not find causality between tourism and the environment in the period between 1995 and 2016 in the 10 countries that attracted the most tourists in 2016. Eyuboglu and Uzar (2020) find bidirectional causality between tourism and the environment in the long run for the period between 1960 and 2014 in Turkey. The analysis results from Koçak et al. (2020) indicate a bidirectional causal relationship between tourism and environmental pollution in the most visited countries. Finally, in Fethi and Senyuçel (2020)'s study, where the authors investigate the role of tourism's development in reducing CO₂ emissions in the top 50 countries that attracted the most tourists, the empirical findings show that the effects of tourism on emission levels are significant and negative, mainly in the cases of France, USA, UK, Germany, Spain, and China, whereas tourism development is positively significant in the cases of Turkey, Thailand, Russia, Greece, Saudi Arabia, Macao, Indonesia, Brazil, Dominica, Philippines, Bulgaria, Tunisia, Egypt, Iran, Georgia, Hong Kong, India, and Malaysia. Jebli et al. (2019), indicate that tourism (particularly green tourism) is a good policy for some countries to combat climate change.

Globalization gained momentum in the Industrial Revolution, and has become a multi-dimensional process today. Whether or not this process is associated with problems in environmental quality has been one of the key areas for discussion. Kalayci and Hayalođlu (2019) found a positive relationship between globalization and environmental pollution in NAFTA countries. The results from Zaidi et al. (2019) on Asia Pacific Economic Cooperation (APEC) countries, point to globalization as the cause of environmental pollution. Likewise, Khan et al. (2019), Khan and Ullah (2019), and Shahbaz et al. (2019) report that globalization increases environmental pollution. The results from Liu et al. (2020) show the validity of an inverted-U relationship, similar to the Environmental Kuznets Curve, between globalization and environmental pollution in G7 countries. Ulucak and Khan (2020), with the data from 1990 to 2015, indicate that globalization has a positive effect on environmental pollution in BRICS countries. However, Saint Akadiri et al. (2019a), Haseeb et al. (2018), Lv and Xu (2018), Shahbaz et al. (2017), and Zafar et al. (2019) find that there is an inverse relationship between the environment and globalization. The results from Saint Akadiri et al. (2020a) show that there is one-way causality from globalization to CO₂ emissions, but according to the generalized impulse response function, CO₂ emissions do not react to changes in globalization.

Multivariate analyses focus on developed countries, Asian countries, the countries that attract the most tourists, emerging economies, and specific groups in terms of the sample (e.g. Brahmašreṇe & Lee, 2017; Saint Akadiri et al., 2019c; Gerceker et al., 2019; Saint Akadiri et al., 2020b). Brahmašreṇe and Lee (2017) explore the relationship between globalization, income, tourism, urbanization, industrialization, and environmental using a panel cointegration test. Their results show that these variables act together, eventually. Saint Akadiri et al. (2019c) analyze the relationship between globalization, income, tourism, energy consumption, and the environment for the 15 countries that attract the most tourists. Their findings show that energy consumption, globalization, and real income increase environmental pollution, while tourism and the square of real income reduce environmental pollution. These results indicate that the Environmental Kuznets Curve hypothesis is valid. Gerceker et al. (2019) investigate the

relationship between globalization, income, tourism, and environment in 16 Mediterranean countries. Their results imply that real income and tourism have a statistically significant effect on the environment. Additionally, increases in real income and tourism have a negative impact on the environment. Saint Akadiri et al. (2020b) investigate the causality between tourism, globalization, economic growth, and CO₂ emissions in 16 small island countries. Their results imply that there is a link from international tourism and economic growth to CO₂ emissions across the panel. However, there is also causality from CO₂ emissions to international tourism.

Saint Akadiri et al. (2019b) and Uzuner et al. (2020) provide an example of time series studies. The sample in both studies is Turkey, but the methods differ. Saint Akadiri et al. (2019b)'s results show that globalization has non-significant negative impacts on CO₂ emissions, while increases in real income levels and international tourists' arrivals cause CO₂ emissions to rise both in the short run and the long run. The causality analysis notes that tourism Granger-causes CO₂ emissions both in the short and long run, while real income and globalization only Granger-cause CO₂ emissions, eventually. Uzuner et al. (2020) only examine causality and their findings show that both the positive and negative shock of tourism influence CO₂ emissions in the long run, while in the short run, only the negative shock of tourism contributes to CO₂ emissions. They also found bidirectional causality between globalization and CO₂ and between economic growth and CO₂.

To investigate the causality between variables in developed countries, Balsalobre-Lorente et al. (2020) focused on OECD countries. According to panel causality test findings, there is a bidirectional causality between globalization, energy, and environmental pollution, and one-way causality from income and tourism to environmental pollution.

3. Data and Measurements

The variables in our study are CO₂ emissions in metric ton per capita, real GDP in US dollars (2010 = 100) (Rgdp), and the tourism variable as the total number of international tourists arriving (Ta). This study was performed according to suggestions in Saint Akadiri et al. (2019b), Akadiri et al. (2020b), Gunduz and Hatemi-J (2005), Katircioglu (2009; 2010; 2014a; 2014b), and Ozcan et al. (2017). The Konjunktur For Schungsstelle (KOF) index from the Swiss Economic Institute developed by Dreher (2006) was used as the globalization measurement unit.

Table 1. Data Description

| Variables | Code | Unit | Source | Period |
|-----------------------|-------------------|-----------------------|--------------------------|-----------|
| Carbon dioxide | lnCO ₂ | Metric ton per capita | WDI | 1995–2014 |
| Economic growth | lnRgdp | Constant (2010 US \$) | WDI | 1995–2014 |
| International tourism | lnTa | Number of arrivals | WDI | 1995–2014 |
| Globalization | lnCof | KOF Index (0–100) | Swiss Economic Institute | 1995–2014 |

The selection of the sample is based on the countries (China, France, Germany, Italy, Mexico, Spain, Thailand, Turkey, the United Kingdom, and the United States) that attract the largest number of tourists (UNWTO, 2019). However, the sample selection is limited by the accessibility of the data. Because of this, it is easier to obtain the possible effects of tourism on the environment in these countries. By choosing variables in this way, we could consider the tourist data from countries, and income and globalization relationships in the context of CO₂ emissions.

In the literature, there are several models to estimate the relationships between these variables. Our study is based on the model (Equation 1) proposed by Saint Akadiri et al. (2019b) for time series and by Gerçeker et al. (2019) for panel data.

$$\ln CO_{2i,t} = \alpha_0 + \beta_1 \ln Rgd p_{i,t} + \beta_2 \ln Ta_{i,t} + \beta_3 \ln cof_{i,t} + \varepsilon_{i,t} \quad (1)$$

Where at cross-section i and period t , $\ln CO_2$ is the natural log of CO_2 emissions, $\ln Rgd p$ is the natural log of real income, $\ln Ta$ is the natural log of tourists arriving, $\ln Cof$ is the natural log of the total globalization index, and ε is the error term.

4. Estimation Methodology

In this study, we use panel unit root tests, panel cointegration tests, and panel cointegration estimators to investigate the effects of globalization, real income, and tourism variables on the environment.

4.1. Panel Unit Root test

Panel unit root tests are divided into two categories according to whether they consider cross-section dependency in groups that form the panel dataset of the test statistics. The tests called first-generation unit root do not consider cross-section dependence, whereas the distinctive feature of the second-generation panel unit root tests is that the test statistics are developed taking into account cross-section dependency (Hurlin and Mignon, 2007).

Many studies depend on panel unit root tests to increase the statistical power of their empirical findings. In this respect, many panel unit root tests are preferred, but we employ unit root tests that complement each other. For instance, LLC is based on the assumption of long-term equilibrium in cross-sections, while IPS assumes that cross-sections are heterogeneous. Additionally, ADF type tests are nonparametric unit root tests. These features were the reason we chose these tests. In this study, the preferred unit root tests, valid for heterogeneous panel samples, are Im, Pesaran and Shin (IPS) (Im et al., 2003), Augmented Dickey-Fuller (ADF), and Phillips-Peron (PP) tests. Levin, Lin and Chu (LLC) (Levin et al., 2002) unit root test is one of the first-generation unit root tests applied to homogeneous panels.

The basic assumptions of the Levin et al. (2002) test are that fixed effects change according to the country, the autocorrelation coefficient is homogeneous for all cross-sections in the panel dataset, and there is no dependency between cross-sections. The LLC test assumes that the autocorrelation coefficient is homogeneous for all cross sections in the panel dataset. Im et al. (2003) developed a test that allowed the autocorrelation coefficient to change, that is, to be heterogeneous, for each unit in the panel. Maddala and Wu (1999) and Choi (2001) proposed a Fisher type test as an alternative, based on combining unit root test statistics for each nonparametric cross-section. For Fisher type test equations, exogenous variables can be determined. External variables cannot be used in test equations, but individual constants (constant effects) or individual constants and trends can be used. Choi (2001) tests the ADF and PP unit root tests, with the null hypothesis and alternative hypothesis with the same hypotheses in the IPS test. Maddala and Wu (1999) are used to determine the number of lags used in ADF regression for each cross-section in the unit root test, while in the Choi (2001) test, Kernel calculation is used as the method for correcting the autocorrelation.

4.2. Panel Cointegration Test

In panel data analyses for testing the long-run cointegration relationship between dependent and independent variables, the panel cointegration tests proposed by Pedroni (1999) are widely used in empirical analyses. In this approach, it is assumed that the dependent and independent variables are not stationary at their levels; therefore, they are stationary in the first difference.

The coefficients of the independent variables in the model of ordinary least square proposed by Pedroni (1999) are variable for each cross-section in the panel. Therefore, the cointegrating vector is heterogeneous for each unit forming the panel. The null hypothesis of this test is “ H_0 : There is no cointegration relationship for all cross sections.”

Pedroni's technique is useful for our study because Pedroni (1999) panel cointegration tests are used to investigate whether the variables are cointegrated in the panel under investigation, given the existence of heterogeneity. This method allows coherent and effective prediction of cointegration vectors and addresses the problem of non-stationary regressors, as well as the problem of simultaneity biases. The statistics from the Pedroni (1999) test give results on seven different cointegration relations. Four of these are known as within-dimension (Panel-V, Panel-RHO, Panel-PP, and Panel-ADF) and three as between-dimension (Group-RHO, Group-PP, and Group-ADF) panel cointegration tests. Between-dimension tests are valid for heterogeneous panels and within-dimension tests are valid for homogeneous panels.

4.3. Panel Cointegration Estimators

We use econometric methods to estimate the in Equation 1 with several coefficient estimators to investigate the theoretical relationships stated in the introduction section. For this purpose, we chose methods that consider heterogeneity with the premise that the slope coefficients are heterogeneous in the long-run coefficient estimates for each panel member.

In this study, we use four different estimators, in two groups. Some of those in the first group are panel Dynamic Ordinary Last Square (DOLS) and Full Modified Ordinary Last Square (FMOLS) predictors developed by Pedroni (2001a, 2001b). These methods estimate the cointegrating vector with dimension data based group mean estimators for each unit and allow the cointegration coefficients to be heterogeneous between the cross-sections. Also, with these methods, panel coefficients can be interpreted as the average of the long run co-integrating vector (Pedroni, 2001b).

The first of the estimators in the second group is the Mean Group (MG), a heterogeneous estimator developed by Pesaran and Smith (1995) for the Panel Autoregressive Distributed Lag (ARDL) model (Blackburne III and Frank, 2007; Pesaran and Smith, 1995). This method considers the heterogeneity of the parameters (Özmen et al., 2019). The main inadequacy of the MG estimator is that it does not allow certain parameters to be equal among the units that make up the panel. This inadequacy in the MG estimator was rectified by the Panel Mean Group (PMG) estimator developed by Pesaran et al. (1999), the second estimator that we have included in this group. With this method, panel long-run or group short-run relationships can be observed. Moreover, the PMG estimator is an approach that accepts that the parameters are (N) heterogeneous between countries in the short and long run.

Panel DOLS and FMOLS estimators are used to solve the problems of internality between the independent variable(s) and error term, as well as auto-correlation in error terms. In addition to this, they are more efficient than the panel OLS. FMOLS and DOLS estimators allow for heterogeneity across individuals and these include individual-specific time trends, individual-specific fixed effects, and time-

specific effects. The estimator is entirely parametric and more precise than the single equation estimators. We also applied MG and PMG for robustness.

5. Empirical Results

In this study, we investigate the impact of globalization, income, and tourism on the environment in the 10 countries that attracted the most tourists in 2019 using panel data techniques. The test and the findings are discussed as follows:

5.1. Results of Unit Root Test

Preliminary tests are important for the consistency and accuracy of the model. Descriptive statistics for the variables included in Equation 1 are presented in Table 2.

Table 2. Descriptive Statistics

| | lnCO₂ | lnRgdp | lnTa | lnCof |
|--------------|-------------------------|---------------|-------------|--------------|
| Mean | 1.838 | 28.222 | 17.246 | 4.298 |
| Median | 1.811 | 28.361 | 17.264 | 4.360 |
| Maximum | 3.004 | 30.418 | 18.242 | 4.491 |
| Minimum | 0.974 | 26.018 | 15.746 | 3.793 |
| Sad. Dev. | 0.524 | 1.035 | 0.612 | 0.156 |
| Skewness | 0.471 | 0.067 | -0.462 | -0.850 |
| Kurtosis | 2.679 | 2.952 | 2.610 | 3.011 |
| Jarque-Bera | 8.251 | 0.171 | 8.398 | 24.123 |
| Probability | 0.016 | 0.917 | 0.015 | 0.0e6 |
| Observations | 200 | 200 | 200 | 200 |

When the results of descriptive statistics are evaluated in terms of distortions, the lnTa and lnCof variables are skewed to the left, lnCO₂ and lnRgdp are skewed to the right. Additionally, lnRgdp is the variable with the highest volatility among the variables. The lnCO₂ and lnCof averages are low compared to lnRgdp and lnTa.

Separate regression results for lnCO₂ and each dependent variable are displayed graphically in Figure 1. In presenting these results this way, it is possible to show that in the literature, these variables were included in the analyses in separate models. These plots regress each variable against all others, and notice the coefficients on each. According to these results, all data points seem to be in range, with no outliers observed.

Table 3 shows a correlation matrix for all variables in the model. Pearson correlation coefficients range from -1 to 1. The closer the value is to 1, the stronger the correlation. The results are presented in this table with the graphic of the correlation matrix (Figure 3). These findings also support the results in Figure 2. The lnRgdp and lnCof correlation matrix coefficients are relatively higher than lnTa.

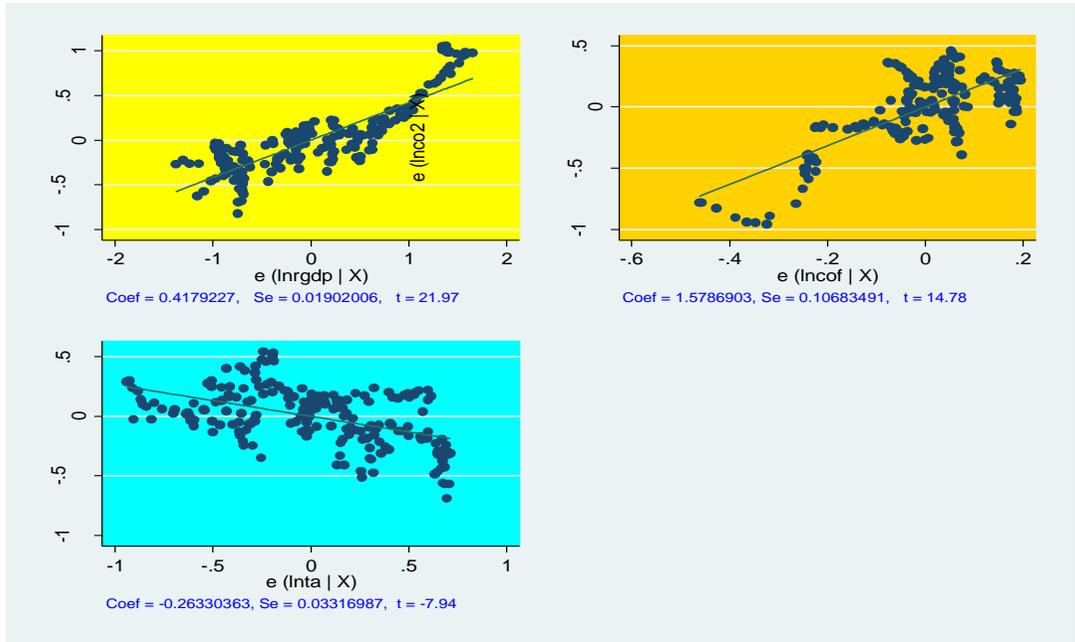


Figure 1. Graphical representation of the regression (lnCO₂; lnRgdp, lnCof, lnTa)

Table 3. Results of Correlation Matrix

| | lnCO ₂ | lnRgdp | lnTa | lnCof |
|-------------------|-------------------|----------------|----------------|-------|
| lnCO ₂ | 1.000 | --- | --- | --- |
| lnRgdp | 0.8188 [0.000] | 1.000 | --- | --- |
| lnTa | 0.4643 [0.000] | 0.6642 [0.000] | 1.000 | --- |
| lnCof | 0.6719 [0.000] | 0.4200 [0.000] | 0.4746 [0.000] | 1.000 |

Notes: The probability selected as 5%. [] indicates the probability values.

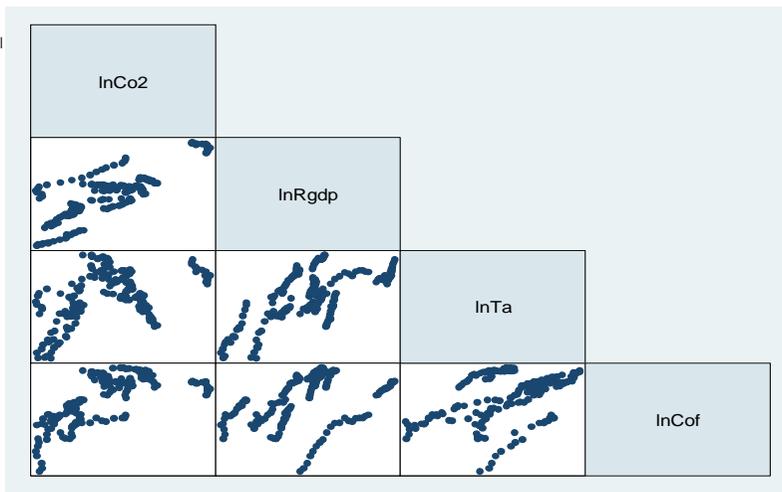


Figure 2. Graphical representation of the correlation matrix

Table 4. Results of Unit Root Tests

| | | Level | | | | First Difference | | | |
|-------------------------|-----|-----------|--------------------|---------------------|--------------------|------------------|--------------------|---------------------|--------------------|
| | | Intercept | | Intercept and Trend | | Intercept | | Intercept and Trend | |
| | | Stat. | prob. | Stat. | prob. | Stat. | prob. | Stat. | prob. |
| lnCO₂ | LLC | 1.528 | 0.936 | -0.691 | 0.244 | -7.605 | 0.000 ^a | -6.101 | 0.000 ^a |
| | IPS | 5.071 | 1.000 | 4.473 | 1.000 | -7.450 | 0.000 ^a | -7.488 | 0.000 ^a |
| | ADF | 7.695 | 0.993 | 23.102 | 0.283 | 93.687 | 0.000 ^a | 87.246 | 0.000 ^a |
| | PP | 6.188 | 0.998 | 15.606 | 0.740 | 118.670 | 0.000 ^a | 121.789 | 0.000 ^a |
| lnRgdp | LLC | -2.276 | 0.011 ^b | -0.674 | 0.250 | -5.687 | 0.000 ^a | -5.827 | 0.000 ^a |
| | IPS | -0.277 | 0.390 | 0.875 | 0.809 | -4.649 | 0.000 ^a | -4.352 | 0.000 ^a |
| | ADF | 26.485 | 0.150 | 20.394 | 0.433 | 57.558 | 0.000 ^a | 53.809 | 0.000 ^a |
| | PP | 33.128 | 0.032 ^c | 13.003 | 0.877 | 69.495 | 0.000 ^a | 75.686 | 0.000 ^a |
| lnTa | LLC | -0.290 | 0.385 | -2.235 | 0.017 ^b | -7.159 | 0.000 ^a | -5.141 | 0.000 ^a |
| | IPS | 2.551 | 0.994 | -1.788 | 0.036 | -7.466 | 0.000 ^a | -6.016 | 0.000 ^a |
| | ADF | 14.135 | 0.823 | 34.146 | 0.025 ^b | 88.361 | 0.000 ^a | 67.875 | 0.000 ^a |
| | PP | 15.885 | 0.723 | 14.053 | 0.827 | 94.042 | 0.000 ^a | 79.046 | 0.000 ^a |
| lnCof | LLC | -8.905 | 0.000 ^a | -1.594 | 0.055 ^c | -8.443 | 0.000 ^a | -8.387 | 0.000 ^a |
| | IPS | -4.4119 | 0.000 ^a | -0.269 | 0.393 | -6.046 | 0.000 ^a | -6.872 | 0.000 ^a |
| | ADF | 54.538 | 0.000 ^a | 24.596 | 0.217 | 71.737 | 0.000 ^a | 76.879 | 0.000 ^a |
| | PP | 78.083 | 0.000 ^a | 22.921 | 0.292 | 66.110 | 0.000 ^a | 69.448 | 0.000 ^a |

Notes: While calculating the long-run consistent error variance in the LLC test, the Barlett method was used as the Kernel estimator, and the bandwidth was selected according to the NeweyWest method.

In the LLC, IPS, ADF, and IPS tests, the maximum lag length was as 3, and the optimal lag length was determined according to the SCI information criterion. a, b, and c show that 1%, 5%, and 10% are statistically significance levels.

According to the unit root test results in Table 4, lnCof is stationary I(0) at level, while it contains unit root at the level according to fixed and trending results, and becomes stationary in the first difference I(1). However, all variables (lnCO₂, lnRgdp, and lnTa) in stationary, trend stationary are I(1) stationary in the first difference.

5.2. Results of Panel Cointegration Test

When it is accepted, considering the trend stationary values of series, that series are I(1) stationary in the first differences, it is important to determine whether they move together in the long run, that is, whether they are co-integrated. The tests developed by Pedroni (1999) to understand this, offer a suitable form for either homogeneous or heterogeneous panel structures. The findings of these cointegration tests are shown in Table 5.

According to Group-PP and Group-ADF that are valid for heterogeneous panels, the series act together eventually for the top 10 tourist destination countries. According to the heterogeneous cointegration test results, CO₂ emissions, economic growth, international tourism, and globalization act together in the long term, and these results are based on strong findings.

Table 5. Results of Cointegration Tests

| Within-dimension test | Constant | | Constant and Trend | |
|---------------------------|----------|-------|--------------------|--------------------|
| | Stat. | prob. | Stat. | prob. |
| Panel-V | -0.749 | 0.773 | 3.800 | 0.000 ^a |
| Panel-RHO | 1.941 | 0.973 | 1.487 | 0.931 |
| Panel-PP (non-parametric) | 1.793 | 0.963 | -1.520 | 0.064 ^c |
| Panel-ADF (parametric) | 1.877 | 0.969 | -3.101 | 0.001 ^a |
| Between-dimension test | Constant | | Constant and Trend | |
| | Stat. | prob. | Stat. | prob. |
| Group-RHO | 2.221 | 0.986 | 2.018 | 0.978 |
| Group-PP (non-parametric) | 0.840 | 0.799 | -5.663 | 0.000 ^a |
| Group-ADF (parametric) | 0.740 | 0.770 | -4.463 | 0.000 ^a |

Notes: In Pedroni tests, the maximum lag length as 1, and the optimal lag length is determined according to the SCI information criterion. Bartlett kernel and Newey-West automatic bandwidth selection were used for non-parametric (PP-type) tests. a, b, and c show that 1%, 5%, and 10% are statistically significance levels.

5.3. Results of Panel Cointegration Estimators

The findings from the FMOLS and DOLS estimators, which we use to estimate the long-run coefficients based on the unit root test results of series stationary in the first difference, are shown in Table 6.

Table 6. Results of FMOLS and DOLS

| Country | FMOLS | | | DOLS | | |
|---------------|------------------------------|-------------------------------|--------------------------------|-------------------------------|-------------------------------|--------------------------------|
| | lnRgdp | lnTa | lnCof | lnRgdp | lnTa | lnCof |
| China | 4.83 [0.000] ^a | 0.52 [0.258] | -1.65 [0.142] | 1.84 [0.000] ^a | -2.38 [0.002] ^a | 1.80 [0.052] ^c |
| France | 1.55 [0.095] ^c | 0.94 [0.157] | -9.66 [0.013] ^b | 2.44 [0.030] ^b | 1.81 [0.127] | -14.22 [0.024] ^b |
| Germany | 0.75 [0.062] ^c | -0.32 [0.001] ^a | -0.86 [0.004] ^c | 3.77 [0.101] | -1.06 [0.074] ^c | -2.93 [0.108] |
| Italy | 1.68 [0.000] ^a | -1.36 [0.000] ^a | 1.07 [0.470] | -0.37 [0.287] | -3.24 [0.000] ^a | 9.32 [0.001] ^a |
| Mexico | 0.34 [0.363] | -0.50 [0.005] ^a | 0.33 [0.617] | 1.58 [0.045] ^b | -0.822 [0.124] | -1.37 [0.227] |
| Spain | 1.18 [0.055] ^c | 1.06 [0.079] ^c | -8.61 [0.006] ^a | 0.95 [0.039] ^b | -0.65 [0.225] | 0.43 [0.834] |
| Thailand | 0.80 [0.000] ^a | -0.04 [0.580] | 0.11 [0.773] | 0.95 [0.000] ^a | 0.16 [0.060] ^c | -1.58 [0.007] ^a |
| Turkey | 0.37 [0.025] ^b | 0.01 [0.953] | 0.46 [0.384] | -0.09 [0.637] | 0.33 [0.033] ^b | -0.60 [0.636] |
| UK | 2.45 [0.077] ^c | -1.09 [0.011] ^b | -10.34 [0.054] ^c | 6.27 [0.000] ^a | -2.28 [0.000] ^a | -21.85 [0.000] ^a |
| United States | -0.36 [0.214] | -0.42 [0.000] ^c | 1.79 [0.215] | -1.10 [0.019] ^b | -0.59 [0.001] ^a | 6.66 [0.035] ^b |
| Panel | 0.89 [0.000] ^a | -0.15 [0.052] ^c | -0.92 [0.076] ^c | 1.13 [0.000] ^b | -0.09 [0.183] | -1.59 [0.014] ^b |

Notes: Bartlett kernel and “Newey-West automatic” bandwidth method were used for FMOLS. In the DOLS estimator, the maximum lag length number is 1 and the optimal lengths are determining according to the SCI information criterion. [] symbolizes probability values. UK: United Kingdom. a, b, and c show that 1%, 5%, and 10% are statistically significance levels.

According to the FMOLS coefficient estimation results for each country in Table 6, throughout the panel, the $\ln R_{gdp}$ variable has positive effects on $\ln CO_2$ emissions whereas the $\ln Ta$ and $\ln Cof$ variables stand out with their negative effects on $\ln CO_2$ emissions. In this context, the $\ln R_{gdp}$ variable has a positive sign in all countries except Mexico and the United States, while the $\ln Ta$ variable is negatively marked in Germany, Italy, Mexico, the UK, and the United States. However, in Spain, the $\ln Ta$ variable has a positive effect on CO_2 emissions. Finally, the $\ln Cof$ variable has a negative effect in France, Germany, Spain, and the UK.

In the DOLS coefficient estimation results for each country, the $\ln R_{gdp}$ variable is positively marked in China, France, Mexico, Spain, Thailand, and the UK. However, $\ln TA$ has a negative effect on $\ln CO_2$ emissions in China, Germany, Italy, the UK, and the United States. $\ln Cof$ is positively marked in China, Italy, and the United States, and negatively marked in France, Thailand, and the UK. It is also worth noting that the coefficients of the $\ln Cof$ variable are relatively high.

According to the findings of the stationary model, within the scope of unit root test results, $\ln Cof$ is stationary $I(0)$ at level, while the trend stationary unit root findings showed that the series did not contain unit root in the first difference of the series. Therefore, when making short-run coefficient estimates, PMG and MG estimators can provide efficient results with the condition that the dependent variable ($\ln CO_2$) is stationary in the first difference and in cases where independent variables are stationary at different levels $I(0)$, $I(1)$ or $I(2)$. Information on which of these tests is more efficient is based on the findings from the Hausman test.

Short-run PMG and MG estimators based on Panel ARDL belonging to each country and the Hausman test findings are shown in Table 7.

In the Hausman test results in Table 7, the null hypothesis of “ H_0 : The difference between MG and PMG estimates is significant” and PMG is more efficient cannot be rejected. Therefore, we can assume that PMG results are more efficient at this stage.

According to the short-run PMG coefficient estimation results for each country shown in Table 6, $\ln R_{gdp}$ (as in FMOLS and DOLS results) stands out with its positive effects. In the PMG results, $\ln R_{gdp}$ coefficients for all countries, except China, are statistically significant and positively marked. The coefficient results of $\ln Ta$ are positive in Spain whereas, in Mexico negatively marked. $\ln Cof$'s country estimate coefficient is statistically significant and negative only in Italy.¹

Finally, when the FMOLS and DOLS panel general coefficient estimation results shown in Table 6 and Table 7 are compared with PMG panel coefficient estimation results, all estimators appear to have similar results regarding the signs of the coefficient estimates for the $\ln R_{gdp}$ (positive) and $\ln Cof$ (negative) variables. On the other hand, while the overall coefficient estimation results of $\ln Ta$ are not statistically significant according to the PMG and DOLS estimators, according to the FMOLS estimator, it is statistically significant and negative.

Especially, when short-run PMG and long-run FMOLS and DOLS results are evaluated together, it can be determined that the EKC hypothesis is valid only in the UK. It can be determined that the results of the $\ln R_{gdp}$ coefficient, which was positive in the short run (PMG) in the UK, turned negative in the long run (DOLS). While in Mexico and Spain, the coefficient of $\ln Ta$ was negative and statistically significant in the short (PMG) and long run (FMOLS). Additionally, the short-run (PMG) and long-run

(FMOLS) results of the $\ln\text{Cof}$ variable in Italy are negative and statistically significant. The findings are summarized in Table 8 for the panel group effects.

Table 7. Results of MG and PMG Short-Run (ARDL 1,1,1)

| Country | MG | | | PMG | | |
|---------------|------------------------------|------------------------------|-------------------------------|------------------------------|-------------------------------|-------------------------------|
| | $\ln\text{Rgdp}$ | $\ln\text{Ta}$ | $\ln\text{Cof}$ | $\ln\text{Rgdp}$ | $\ln\text{Ta}$ | $\ln\text{Cof}$ |
| China | 1.78 [0.094] ^c | 0.17 [0.391] | -1.37 [0.219] | 0.76 [0.267] | -0.04 [0.697] | -0.26 [0.646] |
| France | 1.07 [0.318] | -0.40 [0.406] | -2.41 [0.324] | 1.40 [0.055] ^c | -0.52 [0.146] | -1.37 [0.378] |
| Germany | 0.67 [0.252] | 0.19 [0.555] | -2.49 [0.118] | 0.75 [0.030] ^c | 0.18 [0.353] | -1.77 [0.128] |
| Italy | 1.76 [0.000] ^a | -0.01 [0.889] | -0.90 [0.212] | 1.82 [0.000] ^a | -0.04 [0.624] | -0.97 [0.030] ^b |
| Mexico | 0.43 [0.152] | -0.14 [0.405] | 0.14 [0.855] | 0.44 [0.026] ^b | -0.18 [0.056] ^c | 0.22 [0.668] |
| Spain | 1.21 [0.059] ^c | 0.54 [0.076] ^c | -0.11 [0.942] | 1.19 [0.002] ^a | 0.55 [0.002] ^a | 0.20 [0.859] |
| Thailand | 1.08 [0.000] ^a | -0.19 [0.122] | -0.06 [0.934] | 1.09 [0.000] ^a | -0.13 [0.139] | -0.26 [0.650] |
| Turkey | 0.86 [0.002] ^a | 0.07 [0.935] | 0.05 [0.932] | 0.79 [0.000] ^a | 0.01 [0.795] | 0.05 [0.897] |
| UK | 0.76 [0.415] | -0.07 [0.756] | -1.88 [0.611] | 1.12 [0.062] ^c | -0.10 [0.602] | -0.46 [0.838] |
| United States | 1.07 [0.016] ^b | 0.05 [0.574] | -0.19 [0.828] | 1.07 [0.000] ^a | 0.06 [0.323] | -0.23 [0.723] |
| Panel | 1.07 [0.000] ^a | 0.01 [0.868] | -0.92 [0.005] ^a | 1.04 [0.000] ^a | -0.02 [0.797] | -0.48 [0.023] ^b |

The Hausman test

| | |
|------------|---------|
| Test stat. | 2.80 |
| | [0.247] |

Notes: [] symbolizes probability values. UK: United Kingdom. a, b, and c show that 1%, 5%, and 10% are statistically significance levels.

Table 8. Summary of the Estimator Findings

| Long run | $\ln\text{Rgdp}$ | $\ln\text{Ta}$ | $\ln\text{Cof}$ |
|-----------|------------------|----------------|-----------------|
| FMOLS | + | - | - |
| DOLS | + | - | - |
| Short Run | | | |
| PMG | + | - | - |

6. Discussion and Conclusion

This study tests the impact of globalization, real income, and tourism on the environment in 10 countries that draw the most tourists, using data from the 1995 to 2014 period with panel data techniques. The stability feature was tested with LLC, IPS, ADF, and PP unit root tests and it was found to be stationary in the first differences of the series. Upon this, the tests developed by Pedroni (1999)

were used to determine whether the series act together in the long run. Panel cointegration findings showed that variables act together eventually. In line with the cointegration findings, FMOLS and DOLS techniques for long-run coefficient estimates and PMG and MG results for short-run coefficient estimates are reported.

Our three cointegration estimator results provide the same results both in the long run and in the short run. The income variable has a positive effect on CO₂ emissions in the model. This result has a similar finding to Khan et al. (2020) and Adedoyin et al. (2020), while it has a contrasting finding with Gokmenoglu and Sadeghieh (2019). Similar results from these three studies also show the difficulty of environmentally conscious policies in developing countries.

While Lee and Brahmašreṇe (2013) conclude that tourism reduces environmental pollution, Alam and Paramati (2017), Anser et al. (2020), Sharif et al. (2020b), Jebli et al. (2019), Balsalobre-Lorente et al. (2020), Adedoyina et al. (2020), and Balsalobre-Lorente et al. (2021) find that tourism is the cause of environmental pollution in their sample. The results of this study imply that the tourism variable negatively affects the environment. Fethi and Senyuçel (2020), Gerçeker et al. (2019), Lee and Brahmašreṇe (2013), Paramati et al. (2017b), Saint Akadiri et al. (2019c), and Paramati et al. (2017b) have similar findings. Therefore, it emphasizes that tourism is a green sector.

We see from the literature that the effects of growth and tourism on environmental pollution are relatively less controversial than the effects of globalization. The effect of the globalization variable on the model was found to be negative. Khan and Ullah (2019), Khan et al. (2019), Shahbaz et al. (2019), and Ulucak and Khan (2020) also found a negative effect. In contrast, Kalayci and Hayalođlu (2019) found a positive relationship. In this respect, discussions about whether globalization is positive or negative will continue for a long time. The difference in the findings may be due to sampling and methods. Additionally, the integration of countries into the global economy at different levels may be one of the main factors underlying this. We have also seen that tourism reduces environmental degradation in the sample countries (top 10 countries) in the short and long term. Additionally, the increase in globalization has a negative effect on environmental degradation. These results for tourism countries show that by reducing CO₂ emissions, environmental policies are compatible with macroeconomic targets as an effective tool in the fight against global warming. These findings are consistent with Saint Akadiri et al. (2019c) and it is in line with the points emphasized by Brahmašreṇe and Lee (2017) and Katircioglu (2014, 2014a). In summary, our findings confirm the importance of regulating environmental protection policies, as well as international tourism policies in tourism countries.

According to FMOLS coefficient estimation results, throughout the panel, the increase in real income increases CO₂ emissions, and tourism and globalization decrease CO₂ emissions. In this context, real income negatively affects the environment in all countries except Mexico and the United States. Tourism has a negative impact on the environment in Germany, Italy, Mexico, UK, and the United States. Therefore, these countries should continue their environmental policies to attract more tourists. However, since tourism increases CO₂ emissions in Spain, precautions should be taken to reduce environmental problems in this country. Finally, globalization has a positive effect on the environment in France, Germany, Spain, and the UK. According to the DOLS coefficient estimation results, increases in real income increase environmental pollution in China, France, Mexico, Spain, Thailand, and the UK. Tourism lowers CO₂ emissions in China, Germany, Italy, the UK, and the United States. Globalization increases CO₂ emissions in China, Italy, and the United States while decreasing it in France, Thailand, and the UK.

To determine the short-run effects of globalization, income, and tourism variables, PMG and MG estimators were calculated. However, the Hausman test showed that the PMG estimator was more reliable in determining short-run coefficients. Panel, the signs of the coefficients calculated using the PMG estimator are similar to the long-run results. Thus, while the income variable increases CO₂ emissions in the short-run, the globalization variable decreases CO₂ emissions. These results reveal that the variables of income and globalization have a statistically significant effect on the environment both in the short run and in the long run. However, the direction of causality between each variable and CO₂ emissions is different. As the increase in income stimulates production and consumption activities, it increases the emissions levels. This means these countries should use the increased income also to create a cleaner environment. In the PMG results, real income increases CO₂ emissions in all countries, except China. However, while tourism increases CO₂ emissions in Spain, it decreases them in Mexico. The globalization variable reduces CO₂ emissions only in Italy. Therefore, it is important to use and spread clean technologies to ensure consumption mobility activities and production factors, which increase due to globalization, do not harm environmental balances.

It should not be forgotten that selectivity draws the boundaries of environmental policies in countries. An example of this is the domestic economic dynamics in different countries. Unfavorable economic conditions can tie the hands of policymakers and can hinder sustainable environmental policies. Today, where economic and cultural integration is so intense, regional solutions are limited. However, we can suggest the following policies based on the findings from this study:

- Informative activities should be intensified through written and visual resources to increase environmental awareness in all economic elements.
- The use of renewable energy resources should be increased.
- Clean production techniques should be developed and expanded with some of the income generated from economic growth.
- Natural resources should be used effectively and efficiently.
- In order to generate more income from tourism, efforts to reduce air, water, and soil pollution should be intensified.

There are two main limitations to this study. The first is the sample (i.e. top ten tourist countries) and the second is based on first-generation panel data methods. Additionally, these models can be tested with different econometric methods (e.g. taking into account cross-sectional dependence and structural breaks). We focus only on some econometric methods and variables. For future studies, the variables used in the model can be tested with panel structural break econometric methods. Moreover, the sample can be differentiated and new findings can be compared in terms of country groups. Time series analyzes can be examined to make more comments that are specific about these countries, which could be the subject of future research. In this context, the literature review is useful for further research and the researcher.

Declaration of conflicting interests

The author(s) declared no potential conflicts of interest with respect to the research, authorship, and/or publication of this article.

Notes

1. The model in Equation 1 is named as Double-log model in the literature. Therefore, these results should be interpreted as follows. One per cent change in X ($\ln R_{gdp}/\ln T_a/\ln C_{of}$) will induce a $(B_1/B_2/B_3)$ percent change in Y ($\ln CO_2$).

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